

# Reading Group: Time Series and ML (Week 3)

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**Reading Group:** Time Series and Machine Learning  
(School of Mathematical Sciences)

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- 1 Introduction
- 2 Main Definitions
- 3 Hoeffding's inequality for dependent random variables

## Concentration Inequalities for Dependent Random Variables and Related Studies

- **Article I:** Sara Van De Geer (2002). [On Hoeffding's inequality for dependent random variables](#). Empirical Process Techniques for Dependent Data (Springer Book).
- **Article II:** Kontorovich, L. A., Ramanan, K. (2008). [Concentration inequalities for dependent random variables via the martingale method](#). The Annals of Probability, 36(6), 2126-2158.
- **Article III:** Chang, J., Chen, X., Wu, M. (2022). [Central limit theorems for high dimensional dependent data](#). Forthcoming Bernoulli.

# 1. Introduction

We consider the concept of *pointwise inequalities*, i.e., inequalities that hold uniformly for any  $\theta \in \Theta$ .

Define the function

$$\psi_\alpha(x) := \exp(x^\alpha) - 1, \quad \text{for any } x > 0. \quad (1)$$

For a real-valued random variable  $\xi$ , we define with

$$\|\xi\|_{\psi_\alpha} := \inf \left\{ \lambda > 0 : \mathbb{E} \left[ \psi_\alpha \left( \frac{|\xi|}{\lambda} \right) \right] \leq 1 \right\} \quad (2)$$

Moreover, we write that  $\xi \in \mathcal{L}^q$  for some  $q > 0$  if it holds that

$$\|\xi\|_q := \{\mathbb{E}(|\xi|^q)\}^{1/q} \quad (3)$$

## 2. Main Definitions

### Definition (Orlicz-norm)

For any convex function  $\psi : \mathbb{R}^+ \rightarrow \mathbb{R}^+$  such that  $\psi(0) = 0$  and  $\psi(x) \rightarrow \infty$  as  $x \rightarrow \infty$  and (real-valued) random variable  $X$ , we denote with  $\|x\|_\psi$  the Orlicz-norm, which is defined by

$$\|X\|_\psi := \inf \left\{ C > 0 : \mathbb{E} \left[ \psi \left( \frac{|X|}{C} \right) \right] \leq 1 \right\}. \quad (4)$$

- Denote the  $\ell^p$  Orlicz-norm of  $X$  by  $\|X\|_p$  for  $p \in [0, +\infty)$  by setting  $\psi(x) = x^p$  and  $\|X\|_{e^\gamma}$  the exponential Orlicz-norm for  $\gamma > 0$  by setting  $\psi(x) = \exp(x^\gamma) - 1$  for some  $\gamma \geq 1$ .
- The function  $\psi(x)$  is the convex hull of  $x \mapsto \exp(x^\gamma) - 1$  for some  $\gamma \in (0, 1)$ , which ensures convexity.
- Moreover, when  $\mathbf{X}$  is a random vector, we define its Orlicz-norm by  $\|\mathbf{X}\|_\psi := \sup_{\|\mathbf{u}\| \leq 1} \|\mathbf{u}' \mathbf{X}\|_\psi$ .

## 2. Main Definitions

### Central limit theorems for high dimensional dependent data

Recall that we define with  $S_{n,x} = n^{-1/2} \sum_{t=1}^n X_t$ . Let  $\mathcal{G} \sim \mathcal{N}(0, \Xi)$  where  $\Xi := \text{Cov} \left( n^{-1/2} \sum_{t=1}^n X_t \right)$ . Without loss of generality we assume that  $\mathcal{G}$  is independent of  $\mathcal{X} = \{X_1, \dots, X_n\}$ . We write with  $X_t = (X_{t,1}, \dots, X_{t,p})'$ .

Then, the long-run variance of the  $j$ -th coordinate marginal sequence  $\{X_{t,j}\}_{t=1}^n$  is defined as below

$$V_{n,j} = \text{Var} \left( \frac{1}{\sqrt{n}} \sum_{t=1}^n X_{t,j} \right). \quad (5)$$

Therefore, in order to determine the convergence rate of  $\rho_n$  for the  $\alpha$ -mixing sequence  $\{X_t\}$ , we impose additional regularity conditions.

## 2. Main Definitions

### Assumption (Subexponential moment)

There exists a sequence of constants  $B_n \geq 1$  and a universal constant  $\gamma_1 \geq 1$  such that  $\|X_{t,j}\|_{\psi_{\gamma_1}} \leq B_n$  for all  $t \in [n]$  and  $j \in [p]$ .

### Assumption (Decay of $\alpha$ -mixing coefficients)

There exist some universal constants  $K_1 > 1$ ,  $K_2 > 0$  and  $\gamma_2 > 0$  such that  $\alpha_n(k) \leq K_1 e^{(-K_2 k^{\gamma_2})}$  for any  $k \geq 1$ .

### Assumption (Non-degeneracy)

There exists a universal constant  $K_3 > 0$  such that  $\min_{j \in [p]} V_{n,j} \geq K_3$ .

## 2. Main Definitions

The above condition assumes that the partial sum  $\frac{1}{\sqrt{n}} \sum_{t=1}^n X_{t,j}$  is non-degenerated which is necessary to bound the probability of a Gaussian vector taking values in a small region. When  $\{X_{t,j}\}_{t \geq 1}$  is stationary, then

$$V_{n,j} := \Gamma_j(0) + 2 \sum_{k=1}^{n-1} \left(1 - \frac{k}{n}\right) \Gamma_j(k) \quad (6)$$

where  $\Gamma_k(k) = \text{Cov}(X_{1,j}, X_{k+1,j})$  is the autocovariance of  $\{X_{t,j}\}_{t \geq 1}$  at lag  $k$ .



# Application: Orlicz norm Space

## Example

Consider the inverse of the covariance matrix such that

$$\left| \left| \left| \hat{\Sigma}_i^{-1} \right| \right| \right|_{\psi} \leq C. \quad (7)$$

For the polynomial case, applying the union bound followed by Markov's inequality we conclude that

$$\max_i \left\| \hat{\Sigma}_i^{-1} \right\| \leq_{\mathbb{P}} n^{1/p} \quad \text{and} \quad \max_{i,t,j} \left| X_{i,t}^{(j)} \right| \leq_{\mathbb{P}} (nkT)^{1/p}. \quad (8)$$

## Lemma

*Let  $X$  and  $Y$  be random elements defined in the same probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  taking values in the metric space  $(S, d)$ . Then for measurable  $A$  and  $\delta \geq 0$*

# Application: Orlicz norm Space

Due to the fact that  $x \mapsto \exp \left[ (x / \|X\|_{e^\gamma})^\gamma \right]$  is non-decreasing then

$$\begin{aligned} \mathbb{P}(|X| \geq x) &= \mathbb{P}\left(\exp \left[ (|X| / \|X\|_{e^\gamma})^\gamma \right] \geq \exp \left[ (x / \|X\|_{e^\gamma})^\gamma \right]\right) \\ &\leq \exp \left[ - (x / \|X\|_{e^\gamma})^\gamma \right] \mathbb{E} \exp \left[ (|X| / \|X\|_{e^\gamma})^\gamma \right]. \end{aligned}$$

It holds that,

$$\psi_{e^\gamma}(x) = K_\gamma x \mathbf{1}_{\{0 \leq x \leq a_\gamma\}} + [\exp(x^\gamma) - 1] \mathbf{1}_{\{x \geq a_\gamma\}} \quad (9)$$

where  $K_\gamma := \frac{(\exp a_\gamma^\gamma - 1)}{a_\gamma}$  and  $a_\gamma$  is defined as below

$$a_\gamma := \inf \left\{ x \in \mathbb{R}_+ : x \geq \left( \frac{1 - \gamma}{\gamma} \right)^{1/\gamma} \right\} \quad (10)$$

Moreover, it holds that

$$\left( \frac{1 - \gamma}{\gamma} \right)^{1/\gamma} \leq a_\gamma \leq \left( \frac{1}{\gamma} \right)^{1/\gamma}. \quad (11)$$

### 3. Hoeffding's inequality for dependent random variables

Consider the martingale sequence

$$S_n = \sum_{i=1}^n X_i, n \geq 1. \quad (12)$$

Consider the  $\mathcal{F}_{i-1}$ measurable random variables  $K_i > 0$ , for  $i = 1, 2, \dots$ .  
Define with  $B_0^2 = 0$  and for any  $n \geq 1$  such that

$$B_n^2 = \sum_{i=1}^n K_i^2 \left\{ 1 + \mathbb{E} \left[ \psi \left( \frac{|X_i|}{K_i} \right) \mid \mathcal{F}_{i-1} \right] \right\} \quad (13)$$

### 3. Hoeffding's inequality for dependent random variables

#### Theorem

Let  $\psi$  be an Orlicz function such that it holds that  $\sup_{x,y \rightarrow \infty} \psi(x)\psi(y)/\psi(cxy) < \infty$ , for some constant  $c$ . Suppose that  $\{Z_\theta : \theta \in \Theta\}$  is a separable stochastic process indexed by  $\theta$  in the pseudo-metric space  $(\Theta, \tau)$ . Assume that

$$\|Z_\theta - Z_\vartheta\|_\psi \leq C' \int_0^{\text{diam}(\Theta)} \psi^{-1}(D(\delta)) d\delta \quad (14)$$

where  $\text{diam}(\Theta)$  is the diameter of  $\Theta$  and  $D(\delta)$  is the  $\delta$ -packing number.

### 3. Hoeffding's inequality for dependent random variables

#### Corollary

Let  $W_i$  be  $\mathcal{F}_i$ -measurable and  $\mathbb{E}(W_i|\mathcal{F}_{i-1}) = 0$  for  $i \geq 1$ . Suppose that for some constant  $c < \infty$  it holds that

$$\mathbb{E} \left( \psi \left( \frac{|W_i|}{c} \right) \middle| \mathcal{F}_{i-1} \right) \leq 1, \quad \text{almost surely } i = 1, 2, \dots \quad (15)$$

#### Key Points:

- Partitioning entropy could be applied to nonstationary time series? This could be the case when considering a discretized method, such as block of nonstationary time series (i.e.,  $m$ -dependence).
- Notice that this paper doesn't have in depth explanation of the dependence structure. However, the Orlicz norm provides related moment condition for understanding the asymptotic behaviour.

### 3. Hoeffding's inequality for dependent random variables

- We define with  $\phi(d)$  the following quantity

$$\phi(d) = \int_0^d H^{1/2}(\delta, d) d\delta \vee d := \min \left\{ \int_0^d H^{1/2}(\delta, d) d\delta, d \right\}, \quad (16)$$

- What type of dependence structure does the entropy integral  $\phi(d)$  introduce? For example, what form this integral would have in the case of Garch processes or for the autoregressive model?
- Are there any related results to Hoeffding's inequality for  $\beta$ -mixing sequences? (e.g., **Geometric ergodicity in autoregressive models**)
- To derive the proofs of main results presented in the paper we use that  $P(A) \leq \exp \{-\beta\alpha + 2\beta^2 b^2\}$
- Notice that  $P(A')$  is considered to be a negligible probability.
- All probability bounds are derived with respect to  $S_n$ , which the sum of stationary martingale differences. Similarly, we can consider expressions for partial-sums or partial-sum self normalized processes.

### 3. Hoeffding's inequality for dependent random variables

- Related reference: [Rademacher Complexity of Stationary Sequences](#).
- Define with  $g(y_1, \dots, y_n)$  a measurable function of the data, which for example could be extended to sample moments of estimators.
- Under the assumption of stationary sequences we assume that sub-Gaussianity condition holds in order to obtain probability bounds.
- Furthermore, an important related assumption is the Geometric ergodicity which along with  $\beta$ -mixing can facilitate the development of further the asymptotic theory in time series model. .
- The theoretical framework presented in the paper shows that the theory can be also extended to the case of  $M$  estimators (such as quantile autoregression) using suitable smoothing conditions and deriving the corresponding probability bounds.