

MATH3085/6143 Survival Models – Worksheet 4 Solutions

1. i) The Kolmogorov forward equation has the form

$$\begin{aligned}\frac{d}{dt} {}_t p_x^{23} &= \sum_{k=1}^3 {}_t p_x^{2k} \mu_{x+t}^{k3} && \text{where we define } \mu_{x+t}^{33} = -\mu_{x+t}^{31} - \mu_{x+t}^{32} = 0 \\ &= {}_t p_x^{21} \mu_{x+t}^{13} + {}_t p_x^{22} \mu_{x+t}^{23}\end{aligned}$$

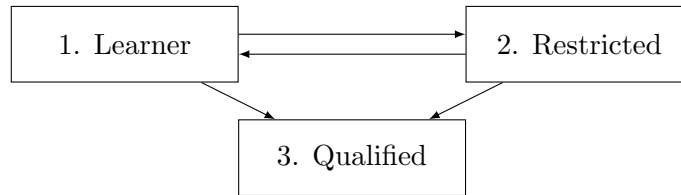
as $\mu_{x+t}^{31} = \mu_{x+t}^{32} = 0$ as state 3 is absorbing.

- ii) The rate of change in the probability of moving from sick to death between time x and $x+t$ is a function of two elements:
- The first term is the rate of death from the healthy state at time $x+t$ weighted by the probability that a person is healthy at $x+t$.
 - The second term is the rate of death from the sick state at time $x+t$ weighted by the probability a person is sick at $x+t$.
- iii) The Kolmogorov forward equation has the form

$$\begin{aligned}\frac{d}{dt} {}_t p_x^{21} &= \sum_{k=1}^3 {}_t p_x^{2k} \mu_{x+t}^{k1} && \text{where we define } \mu_{x+t}^{11} = -\mu_{x+t}^{12} - \mu_{x+t}^{13} \\ &= {}_t p_x^{22} \mu_{x+t}^{21} + {}_t p_x^{23} \mu_{x+t}^{31} - {}_t p_x^{21} (\mu_{x+t}^{12} + \mu_{x+t}^{13}) \\ &= {}_t p_x^{22} \mu_{x+t}^{21} - {}_t p_x^{21} \mu_{x+t}^{12} - {}_t p_x^{21} \mu_{x+t}^{13}\end{aligned}$$

as $\mu_{x+t}^{31} = 0$ as state 3 is absorbing.

2. i)



- ii) Using analogous notations defined in lecture notes, let

- μ_{12} = transition intensity from Learner to Restricted
- μ_{13} = transition intensity from Learner to Qualified
- μ_{21} = transition intensity from Restricted to Learner
- μ_{23} = transition intensity from Restricted to Qualified
- t_1^+ = total observed holding (waiting) time in state Learner
- t_2^+ = total observed holding (waiting) time in state Restricted
- n_{12} = total observed number of transitions from state Learner to state Restricted
- n_{13} = total observed number of transitions from state Learner to state Qualified
- n_{21} = total observed number of transitions from state Restricted to state Learner
- n_{23} = total observed number of transitions from state Restricted to state Qualified.

Therefore, the likelihood function for $\boldsymbol{\mu} = \{\mu_{12}, \mu_{13}, \mu_{21}, \mu_{23}\}$ is given by

$$\begin{aligned}L(\boldsymbol{\mu}) &= \prod_{k,l:k \neq l} \mu_{kl}^{n_{kl}} \exp(-t_k^+ \mu_{kl}) \\ &= \mu_{12}^{n_{12}} \exp(-t_1^+ \mu_{12}) \times \mu_{13}^{n_{13}} \exp(-t_1^+ \mu_{13}) \times \mu_{21}^{n_{21}} \exp(-t_2^+ \mu_{21}) \times \mu_{23}^{n_{23}} \exp(-t_2^+ \mu_{23}) \\ &= \mu_{12}^{n_{12}} \mu_{13}^{n_{13}} \mu_{21}^{n_{21}} \mu_{23}^{n_{23}} \exp[-t_1^+ (\mu_{12} + \mu_{13})] \exp[-t_2^+ (\mu_{21} + \mu_{23})].\end{aligned}$$

iii) (a) It is usually easier to work with log-likelihood,

$$\begin{aligned} l(\boldsymbol{\mu}) &= \log L(\boldsymbol{\mu}) \\ &= n_{12} \log \mu_{12} + n_{13} \log \mu_{13} + n_{21} \log \mu_{21} + n_{23} \log \mu_{23} - t_1^+(\mu_{12} + \mu_{13}) - t_2^+(\mu_{21} + \mu_{23}). \end{aligned}$$

The first order partial derivative of log-likelihood with respect to μ_{12} is then

$$\frac{\partial l(\boldsymbol{\mu})}{\partial \mu_{12}} = \frac{n_{12}}{\mu_{12}} - t_1^+.$$

Now set $\frac{\partial l(\boldsymbol{\mu})}{\partial \mu_{12}} = 0$ and solve for the MLE of the transition rate from Learner to Restricted

$$\frac{n_{12}}{\hat{\mu}_{12}} - t_1^+ = 0 \quad \Rightarrow \quad \hat{\mu}_{12} = \frac{n_{12}}{t_1^+}.$$

Also note that the second order partial derivative is given by

$$\frac{\partial^2 l(\boldsymbol{\mu})}{\partial \mu_{12}^2} = -\frac{n_{12}}{\mu_{12}^2},$$

which is always negative, confirming that the estimate is indeed a maximum.

(b) So from part (iii)(a), the estimated constant transition rate from Learner to Restricted is

$$\hat{\mu}_{12} = \frac{382}{1161} = 0.3290.$$

(c) Recall from lecture notes that the MLEs, $\hat{\mu}_{kl}$ are asymptotically independent, unbiased and normally distributed with $\text{Var}(\hat{\mu}_{kl}) \approx \left[-\frac{\partial^2 l(\boldsymbol{\mu})}{\partial \mu_{kl}^2} \right]^{-1}$, thus we have

$$s.e.(\hat{\mu}_{12}) = \frac{\hat{\mu}_{12}}{n_{12}^{1/2}}.$$

Hence, the 95% confidence interval is given by

$$\hat{\mu}_{12} \pm 1.96 \times s.e.(\hat{\mu}_{12}) = 0.3290 \pm 1.96 \times \frac{0.3290}{382^{1/2}} = [0.296, 0.362].$$

3. i) Using the general expression of maximum likelihood estimate (MLE) of transition intensities in multiple state models, MLE of transition intensity from state 1 to 2, $\hat{\mu}_{12}$ is

$$\hat{\mu}_{12} = \frac{n_{12}}{t_1^+} = \frac{4330}{21650} = 0.2,$$

while the MLE of transition intensity from state 2 to 1, $\hat{\mu}_{21}$ is

$$\hat{\mu}_{21} = \frac{n_{21}}{t_2^+} = \frac{4160}{5200} = 0.8.$$

ii) Denoting $p_{ij}(x, t)$ as the probability that a person in state i at time x will be in state j at time $x + t$ and $\mu_{ij}(x + t)$ ($i \neq j$) as the corresponding transition intensities, then we have from the Kolmogorov forward equations that

$$\begin{aligned} \frac{d}{dt} p_{12}(x, t) &= p_{11}(x, t) \mu_{12}(x + t) - p_{12}(x, t) \mu_{21}(x + t) \\ &= p_{11}(x, t) \mu_{12} - p_{12}(x, t) \mu_{21} \quad (\text{for a time-homogeneous process}) \\ &= [1 - p_{12}(x, t)] \mu_{12} - p_{12}(x, t) \mu_{21} \\ &= \mu_{12} - (\mu_{12} + \mu_{21}) p_{12}(x, t). \end{aligned}$$

Therefore,

$$\begin{aligned}
& \int \frac{dp_{12}(x, t)}{\mu_{12} - [\mu_{12} + \mu_{21}]p_{12}(x, t)} = \int dt \\
\Rightarrow & \frac{\log \{\mu_{12} - [\mu_{12} + \mu_{21}]p_{12}(x, t)\}}{-(\mu_{12} + \mu_{21})} = t + C \\
\Rightarrow & \mu_{12} - [\mu_{12} + \mu_{21}]p_{12}(x, t) = A \exp[-(\mu_{12} + \mu_{21})t] \\
\Rightarrow & p_{12}(x, t) = \frac{\mu_{12} - A \exp[-(\mu_{12} + \mu_{21})t]}{\mu_{12} + \mu_{21}} \\
\Rightarrow & p_{12}(x, t) = \frac{\mu_{12} - \mu_{12} \exp[-(\mu_{12} + \mu_{21})t]}{\mu_{12} + \mu_{21}}
\end{aligned}$$

because we have the boundary condition that $p_{12}(x, 0) = 0$. By substituting the MLE of μ_{12} and μ_{21} obtained from part (i), the required probability is then computed as

$$p_{12}(x, 3) = \frac{0.2 - 0.2 \exp\{-(0.2 + 0.8)3\}}{0.2 + 0.8} = 0.1900.$$

4. i) Similar to question 2(ii), we let

- μ_{12} = transition intensity from state Alive to Dead from heart disease
- μ_{13} = transition intensity from state Alive to Dead from cancer
- μ_{14} = transition intensity from state Alive to Dead from other causes
- t_1^+ = total observed waiting time in state Alive
- n_{12} = total observed number of transitions from state Alive to Dead from heart disease
- n_{13} = total observed number of transitions from state Alive to Dead from cancer
- n_{14} = total observed number of transitions from state Alive to Dead from other causes

The likelihood for the transition intensities, $\boldsymbol{\mu} = \{\mu_{12}, \mu_{13}, \mu_{14}\}$ is given by

$$\begin{aligned}
L(\boldsymbol{\mu}) &= \prod_{k, l: k \neq l} \mu_{kl}^{n_{kl}} \exp(-t_k^+ \mu_{kl}) \\
&= \mu_{12}^{n_{12}} \mu_{13}^{n_{13}} \mu_{14}^{n_{14}} \exp[-t_1^+ (\mu_{12} + \mu_{13} + \mu_{14})].
\end{aligned}$$

ii) The log-likelihood is

$$l(\boldsymbol{\mu}) = \log L(\boldsymbol{\mu}) = n_{12} \log \mu_{12} + n_{13} \log \mu_{13} + n_{14} \log \mu_{14} - t_1^+ (\mu_{12} + \mu_{13} + \mu_{14}).$$

The first order partial derivative of log-likelihood with respect to μ_{12} is then

$$\frac{\partial}{\partial \mu_{12}} l(\boldsymbol{\mu}) = \frac{n_{12}}{\mu_{12}} - t_1^+.$$

Now set $\frac{\partial}{\partial \mu_{12}} l(\boldsymbol{\mu}) = 0$ to obtain the MLE of transition intensity from alive to dead from heart disease as

$$\hat{\mu}_{12} = \frac{n_{12}}{t_1^+}.$$

Also, a simple check shows that the second order partial derivative is

$$\frac{\partial^2}{\partial \mu_{12}^2} l(\boldsymbol{\mu}) = -\frac{n_{12}}{\mu_{12}^2},$$

which is always negative, confirming that the estimate is indeed a maximum.